

## EEL 5544 Noise in Linear Systems Lecture 36

**SPECIAL CASE: WSS GAUSSIAN RPS**

- For Gaussian random processes, the  $n$ -dimensional distribution functions are completely specified by the \_\_\_\_\_ and \_\_\_\_\_
- If  $X(t)$  is Gaussian and WSS, the means and autocovariances are shift-invariant
- Thus, if a random process is Gaussian and WSS, it is \_\_\_\_\_

**MULTIPLE RANDOM PROCESSES**

**DEFN** The *cross-correlation function* for two second-order random processes  $X(t)$  and  $Y(t)$  is

$$R_{XY}(t, s) = E[X(t)Y(s)]$$

**DEFN** The *cross-covariance function* for two second-order random processes  $X(t)$  and  $Y(t)$  is

$$C_{XY}(t, s) = E \{ [X(t) - \mu_X(t)] [Y(s) - \mu_Y(s)] \}$$

Note that

$$C_{XY}(t, s) = R_{XY}(t, s) - \mu_X(t)\mu_Y(s)$$

**DEFN** Two second-order random processes  $X(t)$  and  $Y(t)$  are uncorrelated if

$$C_{XY}(t, s) =$$

$$\Rightarrow R_{XY}(t, s) =$$

**DEFN** Two random processes  $X(t)$  and  $Y(t)$  are *statistically independent* if for each  $n \in \mathbb{Z}^+$  and each  $t_1, t_2, \dots, t_n, s_1, s_2, \dots, s_n \in \mathbb{T}$ ,

$$[X(t_1) X(t_2) \dots X(t_n)] \text{ and } [Y(t_1) Y(t_2) \dots Y(t_n)]$$

are independent random vectors.

- Note that, as with random variables and random vectors, for random processes, s.i.  $\Rightarrow$  uncorrelated, but uncorrelated does not necessarily imply s.i.
- For jointly Gaussian RPs, s.i.  $\Leftrightarrow$  uncorrelated

**DEFN** Two RPs  $X(t)$  and  $Y(t)$  are *jointly wide-sense stationary* if

- **jointly WSS** is a **stronger condition** than each RP being WSS
- for jointly WSS RPs,

$$R_{X,Y}(t, s) = R_{X,Y}(t - s) = R_{X,Y}(\tau).$$

- If  $X(t), Y(s)$  are uncorrelated and  $X(t)$  and  $Y(t)$  are WSS, then

$$C_{X,Y}(t, s) = 0 \Rightarrow R_{X,Y}(t, x) = \mu_X \mu_Y$$

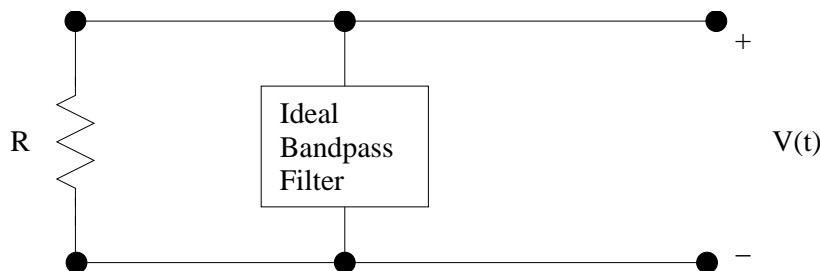
$\Rightarrow R_{X,Y}(t, s)$  is a constant that does not depend on either  $t$  or  $s$

$\therefore$  uncorrelated RPs that are individually WSS are also jointly WSS

### Examples

#### WHITE NOISE

- For second-order RPs,  $E[X^2(t)] < \infty, \forall t \in \mathbf{T}$ .
- **White noise** is a mathematical idealization of the thermal noise process
- Consider thermal noise measurement system:



- measurement shows power  $4kTRB$  in each frequency band  $B$

⇒ infinite power

1. model not valid for extremely high freqs.
2. no system has infinite bandwidth

∴ we assume that thermal noise can be modeled as white noise, with equal power at all frequencies

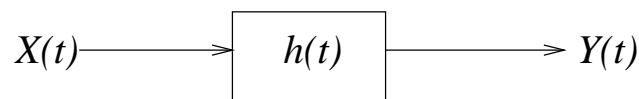
- autocorrelation fcn for white noise is

$$R_X(\tau) = \frac{N_0}{2} \delta(\tau), \quad (1)$$

where the value of  $N_0/2$  is called the (two-sided) noise power spectral density

- We use white Gaussian noise (WGN) to model thermal noise. For engineering purposes, we treat WGN as a zero-mean Gaussian random process with autocorrelation function given by (1)
- If  $X(t)$  is WGN, then  $E[X^2(t)] = R_X(0) = \infty$ , so the mean does not necessarily exist.
- A more careful definition of WGN is  $X(t)$  is a WGN process if, when input to an LTI filter with  $h(t) < \infty$  for all  $t$ , then  $Y(t) = [h * X](t)$  is a zero-mean Gaussian random process.

### TIME-INVARIANT FILTERING



$$\begin{aligned}
 Y &= X * h = h * X \\
 \Rightarrow Y(t) &= \int_{-\infty}^{\infty} X(\tau) h(t - \tau) d\tau
 \end{aligned}$$

- Stability (BIBO = bounded-input, bounded-output):  
Necessary & sufficient condition is  $\exists M_h < \infty \ni$

$$\int_{-\infty}^{\infty} |h(t)| dt \leq M_h.$$

Then  $|X(t)| < \infty \forall t \Rightarrow |Y(t)| < \infty \forall t$ .

CALCULATION OF MEAN AND AUTOCORRELATION

FUNCTIONS FOR WSS RPS

• Suppose:

- $X(t)$  is WSS 2nd-order RP with mean  $\mu_X$  and autocorrelation  $R_X(\tau)$
- $X(t)$  input to BIBO LTI system with impulse response  $h(t)$ .
- $Y(t)$  denotes the output process.

Then

$$\begin{aligned}\mu_Y(t) &= E \left[ \int_{-\infty}^{\infty} X(\tau) h(t - \tau) d\tau \right] \\ &= \int_{-\infty}^{\infty} E[X(\tau)] h(t - \tau) d\tau \\ &= \mu_X \int_{-\infty}^{\infty} h(t - \tau) d\tau \\ &= \mu_X \int_{-\infty}^{\infty} h(\tau) d\tau\end{aligned}$$

$\Rightarrow$  The mean of  $Y(t)$  is a constant.

The autocorrelation function for  $Y(t)$  is

$$\begin{aligned}R_Y(t_1, t_2) &= E \{Y(t_1)Y(t_2)\} \\ &= E \left\{ \int_{-\infty}^{\infty} X(t_1 - u) h(u) du \int_{-\infty}^{\infty} X(t_2 - v) h(v) dv \right\} \\ &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} E \left[ X(t_1 - u) X(t_2 - v) \right] h(u) h(v) du dv \\ &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} R_X(t_1 - t_2 - u + v) h(u) h(v) du dv\end{aligned}$$

- Note that  $R_Y$  only depends on  $t_1$  and  $t_2$  through  $t_1 - t_2$ .

This with the fact that the mean of  $Y(t)$  is constant  $\Rightarrow Y(t)$  is also WSS.

- The expression for  $R_Y$  can be further simplified as

$$\begin{aligned}R_Y(\tau) &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} R_X(\tau - u + v) h(u) h(v) du dv \\ &= (h * \tilde{h} * R_X)(\tau),\end{aligned}\tag{2}$$

where  $\tilde{h}(t) = h(-t)$  (the time-reverse of  $h(t)$ )

- To make computation even simpler, it is often convenient to compute an intermediate function.

Let

$$\begin{aligned} f(\tau) &= (h * \tilde{h})(\tau) \\ &= \int_{-\infty}^{\infty} h(t)h(t - \tau)d\tau. \end{aligned}$$

Then

$$R_Y(\tau) = (f * R_X)(\tau) \quad (3)$$

- Note that:

- $h(t - \tau)$  is a time-shifted version of  $h(t)$
- typically computing  $f(\tau)$  is easier than performing normal convolution
- $f(\tau)$  is an even function

**SPECIAL CASE: : WHITE GAUSSIAN NOISE (WGN)**

- For WGN,  $R_X(\tau) = \frac{N_0}{2}\delta(\tau)$
- Then using (3) and the sifting property of the delta function,  $R_Y(\tau) = \frac{N_0}{2}f(\tau)$